Active Learning in Graph-Based Semi-Supervised Learning

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 $\label{eq:observe} \text{Observe labeled data } \mathcal{D}_\ell = \{(\mathbf{x}_i, y_i)\}_{i \in \mathcal{L}} \text{ and } \textit{unlabeled data } \mathcal{X}_\mathcal{U} = \{\mathbf{x}_j\}_{j \in \mathcal{U}}.$

- $\mathbf{Z} = \{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N\} = \mathcal{X}_{\mathcal{L}} \cup \mathcal{X}_{\mathcal{U}}$
- \mathcal{L} : labeled indices
- \mathcal{U} : unlabeled indices
- $\blacksquare Z = \mathcal{L} \cup \mathcal{U}$

Semi-Supervised Learning

From the given data, can we accurately infer the labelings on U?



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Semi-Supervised Learning

From the given data, can we accurately infer the labelings on \mathcal{U} ?

Active Learning

From the given data, can we judiciously "choose" unlabeled points $\mathcal{Q} \subset \mathcal{U}$ to label that will improve the output of the underlying learning model?

Active Learning vs Semi-Supervised Learning



Active Learning







Given data $\mathcal{X} = \{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N\}$, construct *similarity graph* G(Z, W), where

- $Z = \{1, 2, \dots, N\}$
- $\bullet W_{ij} = \kappa(\mathbf{x}_i, \mathbf{x}_j)$
- $\bullet \ d_i = \sum_{j \in Z} W_{ij}$
- degree matrix $D = \operatorname{diag}(d_1, d_2, \dots, d_N)$

Graph Laplacians

- L = D W, unnormalized
- $L_n = I D^{-1/2} W D^{-1/2}$, normalized
- $L_{rw} = I D^{-1}W$, random walk



Consider family of graph-based SSL models, using a perturbed graph Laplacian $L_{\tau} = L + \tau^2 I:$

$$\hat{\mathbf{u}} = \operatorname*{arg\,min}_{\mathbf{u}\in\mathbb{R}^N} \frac{1}{2} \langle \mathbf{u}, L_{\tau} \mathbf{u} \rangle + \sum_{j\in\mathcal{L}} \ell(u_j, y_j) =: \operatorname*{arg\,min}_{\mathbf{u}\in\mathbb{R}^N} J_{\ell}(\mathbf{u}; \mathbf{y}), \tag{1}$$

for different loss functions ℓ with parameter γ :

$$\begin{split} & \ell(x,y) = (x-y)^2/2\gamma^2, \quad (\text{Regression}) \\ & \ell(x,y) = \ln(1+e^{-xy/\gamma}), \quad (\text{Logistic}) \\ & \ell(x,y) = -\ln\Psi_\gamma(xy), \quad (\text{Probit}) \\ & \text{where } \Psi_\gamma(t) = \int_{-\infty}^t \psi_\gamma(s) ds \text{ is CDF of log-concave PDF } \psi_\gamma(s). \end{split}$$



With perturbed graph Laplacian L_{τ} and n_c the number of classes,

$$\hat{U} = \underset{U \in \mathbb{R}^{N \times n_c}}{\arg\min} \ \frac{1}{2} \langle U, L_{\tau}U \rangle_F + \sum_{j \in \mathcal{L}} \ell(\mathbf{u}^j, \mathbf{y}^j) =: \underset{U \in \mathbb{R}^{N \times n_c}}{\arg\min} \ \mathcal{J}_{\ell}(U; Y),$$

for different loss functions ℓ with parameter γ :

•
$$\ell(\mathbf{s}, \mathbf{t}) = \frac{1}{2\gamma^2} \|\mathbf{s} - \mathbf{t}\|_2^2$$
, (Multiclass Regression)
• $\ell(\mathbf{s}, \mathbf{t}) = -\sum_{c=1}^{n_c} t_c \ln(s_c)$, (Cross-Entropy)

Optimizer $\hat{\mathbf{u}}$ can be viewed as *maximum a posteriori* (MAP) estimator

$$\underset{\mathbf{u}}{\operatorname{arg\,min}} J_{\ell}(\mathbf{u}; \mathbf{y}) \iff \underset{\mathbf{u}}{\operatorname{arg\,max}} \exp(-J_{\ell}(\mathbf{u}; \mathbf{y}))$$

$$= \underset{\mathbf{u}}{\operatorname{arg\,max}} \underbrace{\exp\left(-\frac{1}{2}\langle \mathbf{u}, L_{\tau} \mathbf{u}\rangle\right)}_{prior} \underbrace{\exp\left(-\sum_{j \in \mathcal{L}} \ell(u_j, y_j)\right)}_{likelihood}$$

$$= \underset{\mathbf{u}}{\operatorname{arg\,max}} \mathbb{P}(\mathbf{u}|\mathbf{y})$$

for a posterior distribution $\mathbb{P}(\mathbf{u}|\mathbf{y}) \propto \exp(-J_{\ell}(\mathbf{u};\mathbf{y})).$

Different loss functions give different likelihoods



Harmonic Functions (HF) Model

Assuming hard constraints for labeling¹, have conditional distribution:

$$\mathbf{u}_{\mathcal{U}}|\mathbf{y} \sim \mathcal{N}(\mathbf{u}_{hf}, L_{\mathcal{U},\mathcal{U}}^{-1}), \ \mathbf{u}_{hf} = -L_{\mathcal{U},\mathcal{U}}^{-1}L_{\mathcal{U},\mathcal{L}}\mathbf{y}$$

with $\mathbf{u}_{\mathcal{L}} = \mathbf{y}$.

Gaussian Regression (GR) Model

With $\ell(x,y)=(x-y)^2/2\gamma^2,$ then likelihood/prior/posterior is Gaussian.

$$\mathbb{P}(\mathbf{u}|\mathbf{y}) \propto \exp\left(-\frac{1}{2}\langle \mathbf{u}, L_{\tau}\mathbf{u}\rangle\right) \exp\left(-\frac{1}{2\gamma^2}\sum_{j\in\mathcal{L}}(u_j - y_j)^2\right)$$
$$\sim \mathcal{N}(\mathbf{m}, C), \ \mathbf{m} = \frac{1}{\gamma^2}CP^T\mathbf{y}, \ C^{-1} = L + \frac{1}{\gamma^2}P^TP,$$

where $P: \mathbb{R}^N \to \mathbb{R}^{|\mathcal{L}|}$ is projection onto labeled set \mathcal{L} .

¹Does not actually rigorously fit into Bayesian framework like others



Look-Ahead model with index k and label y_k :

$$\operatorname*{arg\,min}_{\mathbf{u}\in\mathbb{R}^{N}}J^{k}(\mathbf{u};\mathbf{y},y_{k}):=\operatorname*{arg\,min}_{\mathbf{u}\in\mathbb{R}^{N}}\frac{1}{2}\langle\mathbf{u},L_{\tau}\mathbf{u}\rangle+\sum_{j\in\mathcal{L}}\ell(u_{j},y_{j})+\overbrace{\ell(u_{k},y_{k})}^{plus\ k}.$$

 For Gaussian model, look-ahead posterior distribution's parameters from the current posterior distribution

without expensive model retraining – rank-one updates

GR:
$$\mathbf{m}^{k,y_k} = \mathbf{m} + \frac{(y_k - m_k)}{\gamma^2 + C_{kk}} C_{:,k}, \quad C^{k,y_k} = C - \frac{1}{\gamma^2 + C_{kk}} C_{:,k} C_{:,k}^T$$



When likelihood not Gaussian, posterior $\mathbb{P}(\mathbf{u}|\mathbf{y})$ is non-Gaussian.. Problems:

• model classifier as mean $\mu = \mathbb{E}_{\mathbf{u} \sim \mathbb{P}} [\mathbf{u}]$? or MAP estimator $\hat{\mathbf{u}} = \arg \max \mathbb{P}(\mathbf{u} | \mathbf{y})$?

• compute mean, μ , and covariance $C = \mathbb{E}_{\mathbf{u} \sim \mathbb{P}} \left[(\mathbf{u} - \mu)(\mathbf{u} - \mu)^T \right]$? (*potentially expensive!*)

Look-ahead updates??

With non-Gaussian models, we lose these nice properties. What to do?



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Look-ahead updates??

With non-Gaussian models, we lose these nice properties. What to do?

Let's approximate with Gaussian, and see what happens!



Laplace approximation is a popular technique for approximating non-Gaussian distributions \mathbb{P} with a Gaussian distribution.

$$\mathbf{x} \sim \mathcal{N}(\hat{\mathbf{x}}, \hat{C}), \quad \hat{\mathbf{x}} = \operatorname*{arg\,max}_{\mathbf{x} \in \mathbb{R}^N} \ \mathbb{P}(\mathbf{x}), \quad \hat{C} = \left(-\nabla^2 \ln(\mathbb{P}(\mathbf{x}))|_{\mathbf{x} = \hat{\mathbf{x}}}\right)^{-1},$$

where

• $\hat{\mathbf{x}}$: MAP estimator of \mathbb{P}

• \hat{C} : Hessian matrix of the negative-log density of \mathbb{P} , evaluated at $\hat{\mathbf{x}}$



Figure 1: photo credit : http://wiljohn.top/2019/04/14/PRML4-4/



Consider only first M < N eigenvalues and eigenvectors of graph Laplacian, L:



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$$0 = \lambda_1 \leq \lambda_2 \leq \ldots \leq \lambda_M, \quad \mathbf{v}_1, \mathbf{v}_2, \ldots, \mathbf{v}_M.$$

$$\bullet \Lambda_\tau = \operatorname{diag} \left(\lambda_1 + \tau^2, \ldots, \lambda_M + \tau^2\right)$$

$$\bullet V = [\mathbf{v}_1 \quad \mathbf{v}_2 \quad \ldots \quad \mathbf{v}_M] \in \mathbb{R}^{N \times M}$$

$$\bullet \alpha \in \mathbb{R}^M \text{ (binary), } A \in \mathbb{R}^{M \times n_c} \text{ (multiclass)}$$

Binary: $(\mathbf{u} = V\alpha)$
 $J_\ell(\mathbf{u}; \mathbf{y}) = \frac{1}{2} \langle \mathbf{u}, L_\tau \mathbf{u} \rangle + \sum_{j \in \mathcal{L}} \ell(u_j, y_j)$

$$o rac{1}{2} \langle oldsymbollpha, \Lambda_ au oldsymbollpha
angle + \sum_{j \in \mathcal{L}} \ell(\mathbf{e}_j^T V oldsymbollpha, y_j) =: ilde{J}_\ell(oldsymbollpha; \mathbf{y}),$$

Multiclass: (U = VA) $\mathcal{J}_{\ell}(U;Y) = \frac{1}{2} \langle U, L_{\tau}U \rangle_F + \sum_{j \in \mathcal{L}} \ell(\mathbf{u}^j, \mathbf{y}^j)$ $\rightarrow \frac{1}{2} \langle A, \Lambda_{\tau}A \rangle_F + \sum_{j \in \mathcal{L}} \ell(\mathbf{e}_j^T VA, \mathbf{y}^j) =: \tilde{\mathcal{J}}_{\ell}(A;Y).$

$$\boldsymbol{\alpha} | \mathbf{y} \sim \mathcal{N}(\hat{\boldsymbol{\alpha}}, \hat{C}_{\hat{\boldsymbol{\alpha}}}), \qquad \hat{\boldsymbol{\alpha}} = \operatorname*{arg\,min}_{\boldsymbol{\alpha} \in \mathbb{R}^M} \ \tilde{J}_{\ell}(\boldsymbol{\alpha}; \mathbf{y}),$$

and then calculate covariance of Laplace Approximation $\hat{C}_{oldsymbol{lpha}}$

$$\nabla_{\boldsymbol{\alpha}} \tilde{J}_{\ell}(\boldsymbol{\alpha}; \mathbf{y}) = \Lambda_{\tau} \boldsymbol{\alpha} + \sum_{j \in \mathcal{L}} F(\mathbf{e}_{j}^{T} V \boldsymbol{\alpha}, y_{j}) V^{T} \mathbf{e}_{j} = \Lambda_{\tau} \boldsymbol{\alpha} + V^{T} \sum_{j \in \mathcal{L}} F(\mathbf{e}_{j}^{T} V \boldsymbol{\alpha}, y_{j}) \mathbf{e}_{j},$$

$$\nabla_{\boldsymbol{\alpha}}^{2} \tilde{J}_{\ell}(\boldsymbol{\alpha}; \mathbf{y}) = \Lambda_{\tau} + V^{T} \left(\sum_{j \in \mathcal{L}} F'(\mathbf{e}_{j}^{T} V \boldsymbol{\alpha}, y_{j}) \mathbf{e}_{j} \mathbf{e}_{j}^{T} \right) V,$$

$$\implies \hat{C}_{\boldsymbol{\alpha}} = \left(\nabla_{\boldsymbol{\alpha}}^{2} \tilde{J}_{\ell}(\boldsymbol{\alpha}; \mathbf{y}) \right)^{-1} = \left(\Lambda_{\tau} + V^{T} \left(\sum_{j \in \mathcal{L}} F'(\mathbf{e}_{j}^{T} V \boldsymbol{\alpha}, y_{j}) \mathbf{e}_{j} \mathbf{e}_{j}^{T} \right) V \right)^{-1}$$

How to approximate look-ahead model update, $\hat{\pmb{\alpha}}^{k,y_k} = \arg\min \tilde{J}_{\ell}^{k,y_k}$?

• have $\hat{C}_{\hat{\alpha}}$ (i.e. *inverse Hessian*)

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How to approximate look-ahead model update, $\hat{\alpha}^{k,y_k} = \arg \min \tilde{J}_{\ell}^{k,y_k}$? • have $\hat{C}_{\hat{\alpha}}$ (i.e. *inverse Hessian*)

Try one step of Newton's method, starting at $\hat{\alpha}$:

$$\begin{split} \tilde{\boldsymbol{\alpha}}^{k,y_k} &= \hat{\boldsymbol{\alpha}} - \left(\nabla_{\boldsymbol{\alpha}}^2 \tilde{J}_{\ell}^{k,y_k}(\hat{\boldsymbol{\alpha}};\mathbf{y},y_k) \right)^{-1} \left(\nabla_{\boldsymbol{\alpha}} \tilde{J}_{\ell}^{k,y_k}(\hat{\boldsymbol{\alpha}};\mathbf{y},y_k) \right) \\ &= \dots \\ &= \hat{\boldsymbol{\alpha}} - \frac{F((\mathbf{v}^k)^T \hat{\boldsymbol{\alpha}},y_k)}{1 + F'((\mathbf{v}^k)^T \boldsymbol{\alpha},y_k)(\mathbf{v}^k)^T \hat{C}_{\hat{\boldsymbol{\alpha}}} \mathbf{v}^k} \hat{C}_{\hat{\boldsymbol{\alpha}}} \mathbf{v}^k \end{split}$$

where

$$F(x,y) := \frac{\partial \ell}{\partial x}(x,y), \ F'(x,y) := \frac{\partial^2 \ell}{\partial x^2}(x,y).$$

How to approximate look-ahead model update, $\hat{lpha}^{k,y_k} = rgmin \, ilde{J}^{k,y_k}_\ell ?$

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where

$$F(x,y) := \frac{\partial \ell}{\partial x}(x,y), \ F'(x,y) := \frac{\partial^2 \ell}{\partial x^2}(x,y).$$

Simple update!

* GR: this reduces to the exact look-ahead update!



Similar result for multiclass case, but a little lengthy to describe...



$$\tilde{A}^{k,y_k} = \hat{A} - \underbrace{\left(\nabla_A^2 \tilde{\mathcal{J}}^{k,y_k}(\hat{A};Y,\mathbf{y}^k)\right)^{-1} \left(\nabla_A \tilde{\mathcal{J}}^{k,y_k}(\hat{A};Y,\mathbf{y}^k)\right)}_{\text{simplifies to be rank } n_c}$$



Calculating the approximate change in a model (i.e. classifier) from the addition of an index k and associated label y_k has been investigated previously².

²Cai, Zhang, and Zhou, "Maximizing Expected Model Change for Active Learning in Regression", 2013; Karzand and Nowak, "MaxiMin Active Learning in Overparameterized Model Classes", 2020.

Calculating the approximate change in a model (i.e. classifier) from the addition of an index k and associated label y_k has been investigated previously².

Employ approximate update (recalling that $V\alpha = \mathbf{u}$):

$$\begin{aligned} \mathcal{A}(k) &= \min_{y_k \in \{\pm 1\}} \|\hat{\mathbf{u}}^{k,y_k} - \hat{\mathbf{u}}\|_2 \approx \min_{y_k \in \{\pm 1\}} \left\|\tilde{\mathbf{u}}^{k,y_k} - \hat{\mathbf{u}}\right\|_2 = \min_{y_k \in \{\pm 1\}} \left\|\tilde{\boldsymbol{\alpha}}^{k,y_k} - \hat{\boldsymbol{\alpha}}\right\|_2 \\ &= \min_{y_k \in \{\pm 1\}} \left\|\frac{F((\mathbf{v}^k)^T \hat{\boldsymbol{\alpha}}, y_k)}{1 + F'((\mathbf{v}^k)^T \hat{\boldsymbol{\alpha}}, y_k)(\mathbf{v}^k)^T \hat{C}_{\hat{\boldsymbol{\alpha}}} \mathbf{v}^k}\right\| \left\|\hat{C}_{\hat{\boldsymbol{\alpha}}} \mathbf{v}^k\right\|_2 \\ &= \min_{y_k \in \{\pm 1\}} \left\|\frac{F(\hat{u}_k, y_k)}{1 + F'(\hat{u}_k, y_k)(\mathbf{v}^k)^T \hat{C}_{\hat{\boldsymbol{\alpha}}} \mathbf{v}^k}\right\| \left\|\hat{C}_{\hat{\boldsymbol{\alpha}}} \mathbf{v}^k\right\|_2, \end{aligned}$$

²Cai, Zhang, and Zhou, "Maximizing Expected Model Change for Active Learning in Regression", 2013; Karzand and Nowak, "MaxiMin Active Learning in Overparameterized Model Classes", 2020.

Binary Experiments





Checkerboard 2 Dataset. 2,000 points sampled uniformly at random from

$[0,1]^2$.

Graph Construction:

- 10 nearest neighbors, ZP scaling
- M = 50 eigenvalues

Experiments:

- initially label 1 per class
- Sequential
 - 200 active learning iterations, select B = 1 query points at each iteration
- Batch
 - 100 active learning iterations, select B = 5 query points at each iteration

Binary - Checkerboard Results





- **DB-RKHS** "data-based" criterion, RKHS model. (Karzand and Nowak, 2020)
- VOPT, SOPT V-Opt (Ji and Han, 2012), ∑-Opt (Ma et al, 2013)
- UNC Uncertainty Sampling (Settles, 2012)
- RAND Random choices

Multiclass Experiments - HSI





(c) Salinas A



(d) Urban

Graph Construction:

- 15 nearest neighbors, cosine similarity
- Zelnik-Perona scaling
- M = 50 eigenvalues

Experiments:

- initially label 2 per class
- Batch
 - 100 active learning iterations, select B = 5 query points at each iteration
 - MGR (Multiclass Gaussian Regression)
 - CE (Cross-Entropy)



Multiclass GR Results:



Cross-Entropy Results:



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AL in GBSSL



- Adapt this to more useful GBSSL models?
 - Currently only viable for convex loss functions (i.e. for Laplace Approximation)
 - e.g. graph MBO posterior is multimodal, so Laplace approximation meaningful?
- Other active learning criterion that take advantage of the nice model properties we have here?

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